

**Theorem** The exponentiation of a  $N(\mu, \sigma^2)$  random variable is a log normal( $\alpha, \beta$ ) random variable.

**Proof** Let the random variable  $X$  have the normal distribution with probability density function

$$f_X(x) = \frac{1}{\sqrt{2\pi}\sigma} e^{-\frac{1}{2}\left(\frac{x-\mu}{\sigma}\right)^2} \quad -\infty < x < \infty.$$

The transformation  $Y = g(X) = e^X$  is a 1-1 transformation from  $\mathcal{X} = \{x \mid -\infty < x < \infty\}$  to  $\mathcal{Y} = \{y \mid y > 0\}$  with inverse  $X = g^{-1}(Y) = \ln(Y)$  and Jacobian

$$\frac{dX}{dY} = \frac{1}{Y}.$$

Therefore by the transformation technique, the probability density function of  $Y$  is

$$\begin{aligned} f_Y(y) &= f_X(g^{-1}(y)) \left| \frac{dx}{dy} \right| \\ &= \left( \frac{1}{\sqrt{2\pi}\sigma} \right) e^{-\frac{1}{2}\left(\frac{\ln(y)-\mu}{\sigma}\right)^2} \left| \frac{1}{y} \right| \quad y > 0. \end{aligned}$$

Let  $\mu = \ln(\alpha)$  and  $\sigma = \beta$ . Then

$$f_Y(y) = \left( \frac{1}{\sqrt{2\pi y} \beta} \right) e^{-\frac{1}{2}\left(\frac{\ln(y/\alpha)}{\beta}\right)^2} \quad y > 0,$$

which is the probability density function of the log normal distribution.

**APPL Verification:** The APPL statements

```
X := NormalRV(mu, sigma);
g := [[x -> exp(x)], [-infinity, infinity]];
Y := Transform(X, g);
Z := LogNormalRV(mu, sigma);
```

yield identical probability density functions for Y and Z.